# Coffee, tea and decaffeinated coffee in relation to hepatocellular carcinoma in a European population: multicentre, prospective cohort study. 

Bamia, C., Lagiou, P., Jenab, M., Trichopoulou, A., Fedirko, V., Aleksandrova, K., Pischon, T., Overvad, K., Olsen, A., Tjonneland, A., Boutron-Ruault, M.C., Fagherazzi, G., Racine, A., Kuhn, T., Boeing, H., Floegel, A., Benetou, V., Palli, D., Grioni, S., Panico, S, Tumino, R., Vineis, P., Bueno-de-Mesquita, H.B., Dik, V.K., Bhoo-Pathy, N., Uiterwaal, C.S., Weiderpass, E., Lund, E., Quiros, J.R., Zamora-Ros, R., Molina-Montes, E., Chirlaque, M.D., Ardanaz, E., Dorronsoro, M., Lindkvist, B., Wallstroem, P., Nilsson, L.M., Sund, M., Khaw, K.T., Wareham, N., Bradbury, K.E., Travis, R.C., Ferrari, P., Duarte-Salles, T., Stepien, M., Gunter, M., Murphy, N., Riboli, E., Trichopoulos, D.

This is the accepted version of the following article:
Bamia, C., Lagiou, P., Jenab, M., Trichopoulou, A., Fedirko, V., Aleksandrova, K., Pischon, T., Overvad, K., Olsen, A., Tjonneland, A., Boutron-Ruault, M.C., Fagherazzi, G., Racine, A., Kuhn, T., Boeing, H., Floegel, A., Benetou, V., Palli, D., Grioni, S., Panico, S, Tumino, R., Vineis, P., Bueno-de-Mesquita, H.B., Dik, V.K., Bhoo-Pathy, N., Uiterwaal, C.S., Weiderpass, E., Lund, E., Quiros, J.R., Zamora-Ros, R., Molina-Montes, E., Chirlaque, M.D., Ardanaz, E., Dorronsoro, M., Lindkvist, B., Wallstroem, P., Nilsson, L.M., Sund, M., Khaw, K.T., Wareham, N., Bradbury, K.E., Travis, R.C., Ferrari, P., Duarte-Salles, T., Stepien, M., Gunter, M., Murphy, N., Riboli, E., Trichopoulos, D. Coffee, tea and decaffeinated coffee in relation to hepatocellular carcinoma in a European population: multicentre, prospective cohort study.
International Journal of Cancer 136(8): 1899-1908, 2015,
which has been published in final form at http://dx.doi.org/10.1002/ijc. 29214
John Wiley \& Sons, Inc.

## TITLE PAGE

## ARTICLE TYPE: Research Article

TITLE:

Coffee, tea and decaffeinated coffee in relation to hepatocellular carcinoma in a European population: multi-centre, prospective cohort study.

## SHORT TITLE:

Coffee, tea and hepatocellular carcinoma

## CORRESPONDING AUTHOR:

Dr. Christina Bamia
Department of Hygiene, Epidemiology and Medical Statistics
University of Athens, Medical School
75 Mikras Asias st.
11527 Goudi, Athens, Greece
Phone: +30 2107462096
Fax: +30 2107462079
E-mail: cbamia@nut.uoa.gr

## ABBREVIATIONS:

HCC: HepatoCellular Carcinoma

HR: Hazard Ratio

CI: Confidence Intervals

EPIC: European Prospective Investigation into Cancer and nutrition

HBV: Hepatitis B virus

HCV: Hepatitis C Virus

HBsAg: Hepatitis B surface Antigen

LRtest: Likelihood Ratio test

## Keywords:

Hepatocellular carcinoma, liver cancer, coffee, tea, EPIC

## NOVELTY AND IMPACT OF THE WORK:

In the first multi-centre, European cohort study on the topic, intakes of coffee and, to a lesser extent, tea were associated, in a dose-dependent pattern, with lower HCC risk; there was no evidence that these associations were modified by HBV/HCV positivity status or other HCC risk factors. Taking into account that coffee/tea are almost universal exposures, these results may have important implications for high HCC risk subjects.

## AUTHORS:

Christina BAMIA ${ }^{1}$; Pagona LAGIOU ${ }^{1,2,3}$; Mazda JENAB ${ }^{4}$; Antonia TRICHOPOULOU ${ }^{1,5}$; Veronika FEDIRKO ${ }^{4,6}$; Krasimira ALEKSANDROVA ${ }^{7}$; Tobias PISCHON ${ }^{7,8}$; Kim OVERVAD ${ }^{9}$; Anja OLSEN ${ }^{10}$; Anne TJØNNELAND ${ }^{10}$; Marie-Christine BOUTRON-RUAULT ${ }^{11,12,13}$; Guy FAGHERAZZI ${ }^{11,12,13}$; Antoine RACINE ${ }^{11,12,13}$; Tilman KUHN $^{14}$; Heiner BOEING ${ }^{7}$; Anna FLOEGEL ${ }^{7}$; Vasiliki BENETOU ${ }^{1,5}$; Domenico PALLI ${ }^{15}$; Sara GRIONI ${ }^{16}$; Salvatore PANICO ${ }^{17}$; Rosario TUMINO ${ }^{18}$; Paolo VINEIS ${ }^{19,20}$; H.B(as) BUENO-DE-MESQUITA ${ }^{20,21,22}$; Vincent K. DIK $^{22}$; Nirmala BHOO-PATHY ${ }^{23,24}$; Cuno SPM UITERWAAL ${ }^{23}$; Elisabete WEIDERPASS ${ }^{25,26,27,28}$; Eiliv LUND ${ }^{25}$; J.Ramón QUIRÓS ${ }^{29}$; Raul ZAMORA-ROS, 4, 30; Esther MOLINA-MONTES ${ }^{31,32,33}$; Maria-Dolores CHIRLAQUE ${ }^{33,34}$; Eva ARDANAZ ${ }^{33,35}$; Miren DORRONSORO ${ }^{36,37}$; Björn LINDKVIST ${ }^{38}$; Peter WALLSTRÖM ${ }^{39}$; Lena Maria NILSSON ${ }^{40,41}$; Malin SUND ${ }^{42}$; Kay-Tee KHAW ${ }^{43}$; Nick WAREHAM ${ }^{43,44}$; Kathryn E BRADBURY ${ }^{45}$; Ruth C TRAVIS ${ }^{45}$; Pietro FERRARI ${ }^{4}$; Talita DUARTE-SALLES ${ }^{4}$; Magdalena STEPIEN ${ }^{4}$; Marc GUNTER ${ }^{20}$; Neil MURPHY ${ }^{20}$; Elio RIBOLI ${ }^{20}$; Dimitrios TRICHOPOULOS ${ }^{2,3,5}$.

Author Affiliations:
${ }^{1}$ Department of Hygiene, Epidemiology and Medical Statistics, University of Athens Medical School, 11527 Athens, Greece;
${ }^{2}$ Bureau of Epidemiologic Research, Academy of Athens, 11527 Athens, Greece;
${ }^{3}$ Department of Epidemiology, Harvard School of Public Health, MA, 02115 Boston, USA;
${ }^{4}$ International Agency for Research on Cancer (IARC-WHO), 69372 Lyon, France
${ }^{5}$ Hellenic Health Foundation, 115 27, Athens, Greece;
${ }^{6}$ Department of Epidemiology, Rollins School of Public Health, Winship Cancer Institute, Emory University, 30322, Atlanta GA, USA
${ }^{7}$ Department of Epidemiology, German Institute of Human Nutrition, PotsdamRehbruecke, 14558 Nuthetal, Germany
${ }^{8}$ Molecular Epidemiology Group, Max Delbrück Center for Molecular Medicine (MDC) Berlin-Buch, D-13125 Berlin, Germany
${ }^{9}$ Department of Epidemiology, School of Public Health, Aarhus University, DK-8000
Aarhus, Denmark
${ }^{10}$ Danish Cancer Society Research Center, DK-2100 Copenhagen, Denmark
${ }^{11}$ Inserm, Centre for Research in Epidemiology and Population Health, U1018, Institut Gustave Roussy, F-94805, Villejuif, Paris, France
${ }^{12}$ Paris South University, UMRS 1018, F-94805, Villejuif, Paris, France
${ }^{13}$ IGR, F-94805, Villejuif, Paris, France
${ }^{14}$ German Cancer Research Center (DKFZ), 69120 Heidelberg, Germany
${ }^{15}$ Molecular and Nutritional Epidemiology Unit, Cancer Research and Prevention Institute - ISPO, 50139 Florence, Italy
${ }^{16}$ Epidemiology and Prevention Unit, Fondazione IRCCS Istituto Nazionale dei Tumori, 20133 Milan, Italy
${ }^{17}$ Department of clinical and experimental medicine, Federico II University, 80131 Naples, Italy
${ }^{18}$ Cancer Registry and Histopathology Unit, "Civic - M.P.Arezzo" Hospital, ASP 97100 Ragusa, Italy
${ }^{19} \mathrm{HuGeF}$ - Human Genetics Foundation, 10126 Torino, Italy
${ }^{20}$ Department of Epidemiology and Biostatistics, School of Public Health, Faculty of Medicine, Imperial College, W2 1NY London, UK
${ }^{21}$ National Institute for Public Health and the Environment (RIVM), 3720 BA Bilthoven, The Netherlands
${ }^{22}$ Department of Gastroenterology and Hepatology, University Medical Centre Utrecht, 3508 GA Utrecht, The Netherlands
${ }^{23}$ Julius Center for Health Sciences and Primary Care, University Medical Center Utrecht, 3508 GA, Utrecht ,The Netherlands
${ }^{24}$ Julius Centre, Faculty of Medicine, University of Malaya, 50603 Lembah Pantai, Kuala Lumpur, Malaysia
${ }^{25}$ Department of Community Medicine, Faculty of Health Sciences, UiT The Arctic University of Norway, N-9037 Tromsø, Norway
${ }^{26}$ Department of Research, Cancer Registry of Norway,NO-0304 Oslo, Norway,
${ }^{27}$ Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, SE 17177 Stockholm, Sweden
${ }^{28}$ Samfundet Folkhälsan, Helsinki, 00250 Finland
${ }^{29}$ Public Health Directorate, CP 33006 Asturias, Spain
${ }^{30}$ Unit of Nutrition, Environment and Cancer, Catalan Institute of Oncology (ICO), Bellvitge Biomedical Research Institute (IDIBELL), 08907 Barcelona, Spain
${ }^{31}$ Andalusian School of Public Health, 18080 Granada, Spain
${ }^{32}$ Instituto de Investigación Biosanitario de Granada (Granada.bs), 18001 Granada, Spain
${ }^{33}$ CIBER Epidemiology and Public Health CIBERESP, 28029 Madrid, Spain
${ }^{34}$ Epidemiology Department, Murcia Regional Health Council, 30008 Murcia, Spain
${ }^{35}$ Navarre Public Health Institute, 31003 Pamplona, Spain
${ }^{36}$ Public Health Direction, Basque Regional Health Department, 4-20013 Donostia, San Sebastian, Spain
${ }^{37}$ Biodonostia Research Institute CIBER Epidemiology and Public Health CIBERESP, s/n 20014 San Sebastian, Spain
${ }^{38}$ Institute of Medicine, Sahlgrenska Academy, University of Gothenburg SE-40530 Gothenburg, Sweden
${ }^{39}$ Nutrition Epidemiology Research Group, Dept. of Clinical Sciences, Lund University, SE-20502 Malmö, Sweden
${ }^{40}$ Public health and clinical medicine, Nutritional research, Umeå University, SE90185 Umeå, Sweden
${ }^{41}$ Arcum, Arctic Research Centre at Umeå University, SE-90185 Umeå, Sweden
${ }^{42}$ Experimental Surgery, Department of Surgical and Perioperative Sciences/Surgery, Umea University, SE-90185 Umea, Sweden,
${ }^{43}$ University of Cambridge, CB2 0SR Cambridge, UK
${ }^{44}$ MRC Epidemiology Unit, CB2 0QQ Cambridge, UK
${ }^{45}$ Cancer Epidemiology Unit, Nuffield Department of Population Health, University of Oxford, Oxford OX3 7LF, UK


#### Abstract

Inverse associations of coffee and/or tea in relation to hepatocellular carcinoma (HCC) risk have been consistently identified in studies conducted mostly in Asia where consumption patterns of such beverages differ from Europe. In the European Prospective Investigation into Cancer and nutrition (EPIC), we identified 201 HCC cases among 486,799 men/women, after a median follow-up of 11 years. We calculated adjusted hazard ratios (HR) for HCC incidence in relation to quintiles/categories of coffee/tea intakes. We found that increased coffee and tea intakes were consistently associated with lower HCC risk. The inverse associations were substantial, monotonic and statistically significant. Coffee consumers in the highest compared to the lowest quintile had lower HCC risk by $72 \%$ (HR: $0.28 ; 95 \%$ confidence intervals (CI): 0.16 to $0.50, P$-trend $<0.001$ ). The corresponding association of tea with HCC risk was 0.41 ( $95 \% \mathrm{CI}: 0.22$ to $0.78, P$-trend $=0.003$ ). There was no compelling evidence of heterogeneity of these associations across strata of important HCC risk factors, including hepatitis B or hepatitis C status (available in a nested case-control study). The inverse, monotonic associations of coffee intake with HCC were apparent for caffeinated ( $P$-trend $=0.009$ ), but not decaffeinated ( $P$ trend $=0.45$ ) coffee for which, however, data were available for a fraction of subjects. Results from this multi-centre, European cohort study strengthen the existing evidence regarding the inverse association between coffee/tea and HCC risk. Given the apparent lack of heterogeneity of these associations by HCC risk factors and that coffee/tea are universal exposures, our results could have important implications for high HCC risk subjects.


## INTRODUCTION

Hepatocellular carcinoma (HCC) accounts for $85-95 \%$ of primary liver cancer cases. ${ }^{1}$ Several risk factors have been identified, but with respect to diet only alcohol, obesity $^{2}$ and aflatoxin ${ }^{3}$ have been documented. The role of other dietary exposures is still under investigation. ${ }^{4}$

Since 2000, the association of coffee and tea consumption with HCC risk has been investigated in case-control and cohort studies conducted mostly in Asia, where exposure patterns to these beverages differ compared to western countries; two cohort studies on coffee intake were conducted in Europe (Finland) ${ }^{5.6}$ and none in the USA. In a major review, the evidence regarding the indicated associations was considered inconclusive. ${ }^{7}$ Recent meta-analyses, however, pointed to a statistically significant inverse association of coffee intake with primary liver cancer or $\mathrm{HCC},{ }^{8}$ as well as, a suggestive inverse association of tea intake with primary liver cancer. ${ }^{9}$ The earlier studies differed in the extent of adjustment for important HCC risk factors. Moreover, coffee and tea were, in general, considered separately, even though they may act as mutual confounders in their respective associations with HCC risk.

In the context of the European Prospective Investigation into Cancer and nutrition (EPIC), we investigated the associations of coffee (overall, caffeinated, decaffeinated) and tea consumption with HCC risk. This is the first cohort study on the topic undertaken in a multi-ethnic European population.

## METHODS

## Participants

EPIC investigates the role of biological, dietary and lifestyle factors in the aetiology of cancer and other chronic diseases. More than 521,000 apparently healthy men and women, aged 25 to 70 years, were recruited between 1992 and 2000 in 23 centres from 10 European countries (Denmark, France, Germany, Greece, Italy, the Netherlands, Norway, Spain, Sweden, and United Kingdom (UK)). The participants were selected from the general population, except for France (women-members of health insurance plans), Utrecht and Florence (women - breast cancer screening), Naples and Norway (women only), the Oxford cohort (which includes a large proportion of vegetarian volunteers) and subsamples of the Italian and Spanish cohorts (mainly members of blood donor associations). Details of EPIC are given elsewhere. ${ }^{10}$

The study was approved by the Ethical Review Board of the International Agency for Research on Cancer, and by the local Ethics Committees in the participating centres. Participants gave informed consent before enrolment. Procedures were in line with the Helsinki declaration.

## Assessment of diet

Dietary intakes over the previous year were assessed at enrolment through validated centre-specific questionnaires which enquired also about cups of coffee and tea
consumption per day, week or month. ${ }^{11}$ In some centres, additional information on caffeinated and decaffeinated coffee consumption was obtained (Germany, Greece, Italy (except Ragusa and Naples), Netherlands, and UK). Information on tea consumption was not available in Norway. Data on type of tea were not available. Based on the information collected by the questionnaires, beverage intakes were calculated in each center in $\mathrm{ml} /$ day in order to take into account differences in cup sizes by region. Total energy intake was calculated through the EPIC Nutrient-DataBase. ${ }^{12}$

Assessment of lifestyle, somatometry and medical history

At enrolment, standardized questionnaires were used to record sociodemographic, lifestyle and medical history data. Measurements of weight, height and waist/hip circumferences were undertaken for most participants- in France, Norway and Oxford, self-reported weight and height were recorded for the majority of the participants

Follow-up of study population and case ascertainment

Incident HCC cases and vital status were ascertained in most centres through record linkage with cancer and death registries. In France, Germany and Greece a combination of methods was used including health insurance records, pathology registries and active follow-up through mailed questionnaires/telephone interviews. Reported incident cancers were subsequently verified through pathology reports,
medical records and discharge diagnoses. Cancer incidence was coded according to the International Classification of Diseases-Oncology-2. HCC was defined as C22.0.

Censoring dates for cancer incidence for centres using registry data ranged from December 2004 (Asturias (Spain)) to December 2008 (Turin (Italy), Sweden and Norway). For centres using active follow-up, last contact ranged from July 2005 (France) to June 2010 (Heidelberg (Germany)).

From the 521,330 EPIC participants, we excluded 23,818 with prevalent cancer, 4,383 with insufficient follow-up, 78 cases with metastatic cancer or ineligible information on histology, 6,192 with no dietary and 60 with no lifestyle questionnaires.

Eventually, 486,799 participants including 201 incident HCC cases were analysed.

Nested case-control study

Within EPIC, a nested case-control study, using data as of 2006 on 125 HCC cases and 250 matched controls was conducted. ${ }^{13}$ For these, serum levels of hepatitis B surface antigen (HBsAg) and antibodies to hepatitis C virus (anti-HCV) were determined at the Centre de Biologie République laboratory in Lyon, France. Design and methods of this study have been described elsewhere. ${ }^{13}$ Controls were selected from those alive and free of cancer at the time of diagnosis of the case. Matching criteria included study centre, sex, age at the time of blood collection, date and time of day of blood collection. Women were additionally matched by menopausal status and by use of exogenous hormones.

## Statistical analysis

The associations of coffee, caffeinated/decaffeinated coffee, and tea intake with HCC risk were investigated through Cox proportional hazard models. Age at exit was the underlying time variable defined as the age of: HCC diagnosis (case); diagnosis of any other incident cancer (censored); death for participants who died without cancer (censored); last follow-up for those alive and with no cancer (censored). Age at recruitment was the entry time. All models were stratified by recruitment age (1-year intervals) and study centre.

Coffee intake was considered in sex-specific quintiles (Q1-Q5). For tea, due to the large number of non-consumers ( $153,097 / 450,921,34 \%$ ), five categories of intake (C1-C5) were created: non-consumers (C1) and the sex-specific quartiles (C2-C5) among consumers. Both variables were considered categorically and per one-level increments. Norway was excluded from tea analyses, as no relevant data were available. Due to potential between-countries variations in coffee/tea volume and concentration, we also based estimation of quintiles of coffee and categories of tea on sex-and-country specific intakes in order to check the robustness of our findings. Linear trends were evaluated using the medians of quintiles/categories as continuous variables in the Cox models. In additional models, coffee and tea were mutually adjusted for.

In analyses by type of coffee, only data from 163,672 participants (Germany, Netherlands and United Kingdom), for whom the sum of caffeinated and decaffeinated coffee equals the total amount of coffee intake, were included. Due to the large number of non-consumers of decaffeinated coffee (35\%) we considered 3 categories for this beverage: (i) C 1 : non-consumers, (ii) C2: consumers below the median (among consumers), and (iii) C3: consumers at/above the median (among consumers). To preserve comparability here, caffeinated coffee and tea were also analysed in 3 categories defined as for decaffeinated coffee.

In all analyses, potential confounders were: sex, self-reported diabetes (categorically: no, yes, unknown), education (categorically: no formal/primary school, technical/secondary school, university /longer, unknown), physical activity (categorically: inactive/moderately inactive, moderately active/active, unknown), smoking (categorically: never, former smoker, current smoker, unknown), ethanol intake (categorically: low: men $<10 \mathrm{~g} /$ day, women $<5 \mathrm{~g} /$ day; moderate: men $10-<40$ $\mathrm{g} /$ day, women $5-<20 \mathrm{~g} /$ day; high: men $\geq 40 \mathrm{~g} /$ day, women $\geq 20 \mathrm{~g} /$ day ), body mass index (categorically: $\leq 25 \mathrm{~kg} / \mathrm{m}^{2},>25$ to $<30 \mathrm{~kg} / \mathrm{m}^{2}, \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ) and energy intake (ordered, in sex-specific quintiles). In order to minimize residual confounding by smoking/ethanol intake, we further used more detailed variables (but with more missing values), accounting for intensity and/or duration of smoking and lifetime alcohol drinking. Deviations from log linearity of the associations between coffee/tea intakes and HCC risk were assessed by: a) comparing models including coffee/tea intakes as ordered, with models including both dietary exposures as categorical, using the Likelihood Ratio test (LRtest), b) by applying restricted cubic splines Cox regression models for coffee/tea intakes and testing (using the Wald test and the

LRtest) whether these models were of better fit as compared to models which assumed simpler, log linear associations between coffee/tea and HCC risk, and c) by inspecting the Akaike's Information Criterion (AIC) in models assuming restricted cubic splines for the indicated associations as compared to models of linear trends which are shown in Table 3.

We performed several sensitivity analyses (see Results). Furthermore, we investigated whether the estimated associations differed across strata defined by sex (men/women), diabetes (yes/no), smoking (never/former/current), ethanol intake (low/moderate/high as indicated above) and body mass index ( $\leq 25 \mathrm{~kg} / \mathrm{m}^{2} />25$ to $<30$ $\mathrm{kg} / \mathrm{m}^{2} / \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ), by introducing interaction terms between quintiles/categories of coffee/tea intake (ordered) and each of the indicated variables (categorically). In order to evaluate whether the associations between coffee/tea intake and HCC risk varied by positivity status regarding hepatitis B virus (HBV) and/or hepatitis C virus (HCV), we used data from the nested case-control study ${ }^{13}$ and introduced interaction terms in conditional logistic regression models between HBV/HCV status (yes/no) and coffee/tea intakes; in this analysis, coffee/tea intakes were considered in tertiles (T1-T3, using control-defined cut-off values) due to the small number of controls. $\mathrm{P}<0.05$ was considered statistically significant. All analyses were performed using STATA (Stata Corporation: Stata statistical software, release 11. College Station. TX: Stata Corporation; 2009).

## RESULTS

After a median follow-up of 11 years, $1,622,408$ and $3,744,060$ person-years were accrued from the $145,039(30 \%)$ men and $341,760(70 \%)$ women of the study population, respectively, and 133 men and 68 women were diagnosed with HCC.

Table 1 presents the distribution of participants according to sex-specific quintiles of coffee intake, country and baseline characteristics. Overall median coffee consumption was $354 \mathrm{ml} / \mathrm{d}$ among men and $290 \mathrm{ml} /$ d among women; $7 \%$ of the participants did not consume coffee. Coffee intake varied across countries, with the Danes consuming the highest quantities. Participants with higher coffee intakes were slightly older, more educated, reported less frequently diagnosis with diabetes, were more frequently smokers, consumed higher amounts of ethanol, had lower body mass index, were more active and had higher energy intake.

In Table 2, corresponding distributions as in Table 1 are presented by categories of tea intake. Among tea consumers ( $66 \%$ of the total cohort), median intakes were 107 $\mathrm{ml} / \mathrm{d}$ for men and $190 \mathrm{ml} / \mathrm{d}$ for women. Distribution of participants across categories of tea intake varied across countries, with the highest intakes consumed in the UK. Consumers of high tea intakes were more educated, more active, former smokers, moderate ethanol drinkers, reported less frequently diabetes, and had lower body mass index.

Table 3 shows hazard ratios (HRs) for HCC incidence by quintiles/categories of coffee/tea intakes, separately (models 1 and 2 ) and after mutual adjustment (model 3).

Increased coffee intake was associated with a substantial, monotonic and statistically significant reduction in HCC risk, in all models. This inverse association was more pronounced when coffee and tea were considered simultaneously $\left(\mathrm{HR}_{\mathrm{Q5vQ}}, 0.28\right.$; $95 \%$ confidence interval (CI), 0.16 to 0.50 ). Tea intake was also inversely associated with HCC incidence, with a linear trend apparent in all models (particularly model 3).

There was no statistical evidence indicating that models assuming a log linear association between coffee/tea intakes and HCC risk were of inferior fit as compared to models with no assumption for the type of association of coffee/tea with HCC risk (i.e., using coffee/tea intakes in quintiles/categories, categorically): LRtest $_{\text {coffee }}$ $=0.707$, LRtest $_{\text {tea }}=0.683$. Similarly, there was no evidence that restricted cubic splines regression models were of better fit as compared to simpler models assuming log linear associations between coffee/tea and HCC risk (p-value from Wald test: coffee: 0.13 ; tea: 0.78 ). The AIC for model 3 of Table 3 of linear trends (AIC $=2372.795$ ) indicated a slightly better model fit as compared to the restricted cubic splines models for coffee (AIC=2379.004) and tea (AIC=2409.564).

More detailed control for smoking (i.e., smoking intensity/duration) and alcohol intake (e.g. lifetime alcohol consumption), as well as, the inclusion of body mass index and energy intake as continuous rather than as categorical variables, generated very similar results.

Sensitivity analyses

Sensitivity analyses focussed on model 3 (Table 3). When subjects without missing values were studied ( 187 HCC cases/413,666 participants), HRs and 95\% CIs were
essentially identical to those reported in Table 3. When 9,596 subjects (10 HCC cases) in the top/bottom $1 \%$ of the energy intake distribution were excluded, results were also similar to those previously reported (coffee: $\mathrm{HR}_{\mathrm{Q} 5 \mathrm{vs} \mathrm{Q}}, 0.26 ; 95 \% \mathrm{CI}, 0.14$ to 0.48 ( $P$-trend $<0.001$ ); tea: $\mathrm{HR}_{\mathrm{C} 5 \mathrm{vCl}}, 0.36 ; 95 \% \mathrm{CI}, 0.18$ to 0.73 ( $P$-trend $=0.002$ )). When analyses were restricted to the 167 histologically confirmed HCC cases with tea consumption information, the inverse associations of coffee/tea with HCC risk followed the same linear trends (coffee: $P$-trend $<0.001$; tea $P$-trend $=0.009$ ). Excluding the first 2 years of follow-up ( 171 HCC cases/442,961 participants analysed) did not alter the strength/pattern of the estimated associations (coffee: $P$ trend $<0.001$; tea: $P$-trend $=0.002$ ). Finally, using sex-and-country specific quintiles/categories for coffee and tea, resulted in somewhat less striking, albeit still statistically significant associations for coffee ( $\mathrm{HR}_{\mathrm{Q} 2 \mathrm{vs} \mathrm{Q}}, 0.57 ; 95 \% \mathrm{CI}, 0.39$ to 0.85 ; $\operatorname{HR}_{\mathrm{Q} 3 v \mathrm{Q} Q}, 0.57 ; 95 \% \mathrm{CI}, 0.36$ to $0.89 ; \mathrm{HR}_{\mathrm{Q} 4 v \mathrm{~s} Q 1}, 0.56 ; 95 \% \mathrm{CI}, 0.36$ to 0.87 ; $\mathrm{HR}_{\mathrm{Q} 5 v s Q 1}$, $0.45 ; 95 \% \mathrm{CI}, 0.28$ to $0.74 ; P$-trend= 0.002 ). For tea, the apparent linear trend was no longer statistically significant, possibly on account of scarcity of data ( $P$-trend $=0.08$ ).

## Examination of effect-modifications

Interactions were evaluated among participants without missing data and were, generally, far from statistically significant. The associations of both beverages were considerably weaker among diabetics, but the interaction terms were statistically nonsignificant (coffee, $P$-interaction $=0.14$; tea, $P$-interaction $=0.20$ ). There was a suggestion for tea-by-sex interaction ( $P$-interaction=$=0.06$ ), possibly on account of the small number of female HCC cases; HRs ( $95 \%$ CIs) per category increase in tea
consumption were 0.80 ( 0.68 to 0.94 ) among men, but 1.08 ( 0.84 to 1.38 ) among women.

In the nested case-control analysis, odds ratios (OR) for HCC risk, adjusted for HBV/HCV status, smoking, ethanol intake, education, and mutually for coffee and tea intakes were: coffee: $\mathrm{OR}_{\mathrm{T} 3 \mathrm{vsT1}}, 0.54$ ( $95 \% \mathrm{CI}, 0.19$ to 1.58 ), $P$-trend $=0.25$; tea: $\mathrm{OR}_{\mathrm{T} 3 \mathrm{vsT} 1}, 0.36(95 \% \mathrm{CI}, 0.13$ to 0.99$), P$-trend 0.04 . No evidence for effect modification of the impact of coffee/tea intakes on HCC risk according to HBV/HCV status was observed (coffee: $P$-interaction $=0.78$; tea: $P$-interaction $=0.50$ ).

Caffeinated and decaffeinated coffee

In analyses of 163,672 participants ( 64 HCC cases) from Germany, Netherlands and UK, the inverse association of coffee with HCC risk was seen across the three categories of caffeinated $(P$-trend $=0.009)$, but not of decaffeinated coffee $(P$ trend $=0.45$ ). Compared to caffeinated coffee, the inverse association of decaffeinated coffee with HCC risk was considerably weaker and statistically non-significant $\left(\mathrm{HR}_{\mathrm{C} 2 \text { vs.C1 }}: 0.94,95 \% \mathrm{CI}, 0.39\right.$ to $2.28 ; \mathrm{HR}_{\mathrm{C} 3 v \mathrm{~s} . \mathrm{Cl}}: 0.74,95 \% \mathrm{CI}, 0.36$ to 1.51 ).

## DISCUSSION

In a multi-centre European cohort, we found evidence that intakes of coffee and, to a lesser extent, tea are associated with lower HCC incidence. Both associations followed statistically significant linear trends, which were also generally apparent in
subgroup analyses. Compared to those with minimal or no consumption of coffee, and adjusted also for tea intake, those in the highest quintile of coffee intake had a statistically significantly lower HCC risk by about $70 \%$. For tea, the corresponding contrast was milder than that for coffee, with the risk in the highest quintile being about $60 \%$ lower in comparison to the lowest quintile (composed exclusively by nonconsumers). This is the first multi-centre cohort study in a western population which has investigated simultaneously and with mutual adjustment the association of coffee and tea intakes with HCC risk.

Our results are in line with those reported in previous case-control ${ }^{14-16}$ and cohort ${ }^{17-22}$ Asian studies, as well as in three case-control studies ${ }^{23-25}$ and in two, moderatelysized, cohort studies conducted in Europe ${ }^{5,6}$ with respect to: a) the inverse association of coffee consumption with HCC risk, b) the linear trend in risk reduction with increasing coffee intake, and c) the consistency of these findings across various subgroups. In a recent meta-analysis, ${ }^{8}$ including all indicated studies, as well as three studies reporting statistically non-significant inverse ${ }^{19,26}$ or null ${ }^{27}$ associations, the summary relative risk for HCC for any vs. no coffee consumption was $0.60(95 \% \mathrm{CI}$, $0.50-0.71$ ), with a dose-response trend. With the exception of one Italian casecontrol study, ${ }^{25}$ studies on tea consumption and HCC risk have been conducted exclusively in Asia (where green tea consumption is more frequent than in Europe); the results are collectively inconclusive. Two studies reported statistically significant inverse associations, ${ }^{28,29}$ four inverse but statistically non-significant associations, ${ }^{30-33}$ two null associations, ${ }^{19,34}$ whereas three reported positive, albeit statistically nonsignificant associations. ${ }^{17,25,35}$ A meta-analysis of these studies ${ }^{9}$ pointed to a suggestive inverse association of tea with primary liver cancer. Our study, the first
cohort study outside Asia, indicates an inverse, dose-dependent association of tea with HCC risk, which is not incompatible with the results of the indicated meta-analysis, although our findings point to a stronger inverse association.

Several mechanisms could underlie the inverse association between coffee intake and HCC. Coffee has been inversely related to chronic liver disease, ${ }^{36,37}$ cirrhosis, ${ }^{37,38}$ or diabetes mellitus. ${ }^{39}$ In previous studies, however, the inverse association between coffee and HCC risk was evident in subjects irrespectively of chronic liver disease, or HBV/HCV status. ${ }^{8}$ When we excluded the first 2 years of follow-up, which would exclude a fraction of cases with underlying liver disease, results were largely unchanged. Moreover, adjusting for chronic HBV/HCV infection in our nested casecontrol study, did not materially affect the results, nor was there a statistically significant interaction between virus carrier state and coffee (and tea) intake with respect to HCC risk - a finding with considerable clinical importance, ${ }^{40}$ given the high HCC risk of $\mathrm{HBV} / \mathrm{HCV}$ carriers. Of note, our results were adjusted for diabetes and there was no evidence for diabetes-by-coffee (or tea) interaction.

Coffee contains several hundred chemical compounds including caffeine, diterpenes, chlorogenic acids and flavonoids, with reported anti-hepatocarginogenic effects. ${ }^{41}$ Regarding caffeine intake, in a subset of our data, intake of caffeinated coffee, tea and decaffeinated coffee appear to be inversely related to HCC risk but the association is strong and statistically significant for caffeinated coffee, weaker and statistically nonsignificant for tea and still weaker for decaffeinated coffee. Taking into account, the correspondence with the caffeine content of the three beverages and given that caffeine content (not content of other antioxidant compounds) ${ }^{42}$ is the defining
difference between caffeinated and decaffeinated coffee, a key role of caffeine cannot be excluded. We did not have available information on the amount of caffeine intake (from all food/beverage sources) in order to better explore the potential role of caffeine on the inverse association of coffee with HCC risk. The possible differential association of caffeinated and decaffeinated coffee with HCC risk needs to be further explored, preferably in other large cohort studies.

Regarding tea, experimental studies have suggested a hepatoprotective effect, mainly invoking antioxidative properties of green tea polyphenols. ${ }^{43}$ There was no adequate information on type of tea in our cohort, but Europeans are likely to consume mostly black rather than green tea. Given the inconclusive collective evidence in the literature, further studies are needed to affirm any protective role of tea on HCC risk before exploring the underlying mechanisms.

Strengths of our study are its cohort design, its multi-centre coverage allowing to examine a variable range of tea and coffee intakes across European countries, the use of validated questionnaires, the relatively long follow-up, the common identification criteria of incident HCC and the available information about potential confounders. A limitation of our study is the unavoidable modest number of HCC outcomes, given the low incidence of the disease in most European countries. Another limitation stems from the lack of data about brewing methods across countries, as well as from complications arising from the different volume/concentration of coffee consumption across countries. Errors in reporting beverage consumption cannot be ruled out, but this is more likely to result in random misclassification of intakes, thus generating underestimation of true associations. Bias in the estimated associations could have
been introduced if cases had modified their coffee/tea consumption during the prediagnostic period of their disease, but excluding cases diagnosed in the first 2years of follow-up generated essentially identical results.

In conclusion, in a large prospective study across European countries, we found evidence that increased coffee and tea intakes are associated, in a dose-dependent pattern, with lower HCC risk. The estimated risk reduction was stronger for coffee than for tea, whereas there was no convincing evidence for an association of decaffeinated coffee with HCC. Given the apparent lack of heterogeneity of the HCC associations with coffee and tea by $\mathrm{HBV} / \mathrm{HCV}$ positivity status or other risk factors, the results may even have important implications for high HCC risk subjects. Mechanistic studies could clarify whether the noted associations could be attributed to caffeine, or, to other constituents of these beverages.

## Acknowledgments:

Reagents for the hepatitis infection determinations were kindly provided by Abbott, France.

## Funding:

This work was supported by the French National Cancer Institute (Institut National du Cancer; INCA), Grant number 2009-139.

The coordination of EPIC is financially supported by the European Commission (DGSANCO) and the International Agency for Research on Cancer (coordination of EPIC). The national cohorts are supported by Danish Cancer Society (Denmark); Ligue Contre le Cancer, Institut Gustave Roussy, Mutuelle Générale de l'Education Nationale, Institut National de la Santé et de la Recherche Médicale (INSERM) (France); Deutsche Krebshilfe, Deutsches Krebsforschungszentrum and Federal Ministry of Education and Research (Germany); Stavros Niarchos Foundation and Hellenic Health Foundation (Greece); Italian Association for Research on Cancer (AIRC), National Research Council, HuGeF Foundation, Compagnia di San Paolo, Sicily Regional Government, AIRE ONLUS Ragusa and AVIS Ragusa (Italy); Dutch Ministry of Public Health, Welfare and Sports (VWS), Netherlands Cancer Registry (NKR), LK Research Funds, Dutch Prevention Funds, Dutch ZON (Zorg Onderzoek Nederland), World Cancer Research Fund (WCRF) and Statistics Netherlands (The Netherlands); ERC-2009-AdG 232997 and Nordforsk, Nordic Centre of Excellence programme on Food, Nutrition and Health, Norwegian Research Council, Norwegian Cancer Society, University of Tromso (Norway); Health Research Fund (FIS), Regional Governments of Andalucía, Asturias, Basque Country, Murcia (no. 6236) and Navarra, ISCIII RETIC (RD06/0020) (Spain); Swedish Cancer Society, Swedish

Scientific Council and Regional Government of Skåne and Västerbotten (Sweden); Cancer Research UK, Medical Research Council (United Kingdom).

None of the funding sources had any role in the design of the study; collection, analysis and interpretation of the data; writing of the report, and decision to submit the article for publication.

## Conflict of interest

The authors declared no conflict of interest.

## REFERENCES

1. El-Serag HB, Rudolph KL.Hepatocellular carcinoma: epidemiology and molecular carcinogenesis. Gastroenterology 2007; 132: 2557-76.
2. Schlesinger S, Aleksandrova K, Pischon T, Fedirko V, Jenab M, Trepo E, Boffetta P, Dahm CC, Overvad K, Tjønneland A, Halkjær J, Fagherazzi G, et al. Abdominal obesity, weight gain during adulthood and risk of liver and biliary tract cancer in a European cohort. Int J Cancer 2013;132: 645-57.
3. International Agency for Research on Cancer. Aflatoxins. In: Traditional Herbal Medicines, Some Mycotoxins, Naphthalene and Styrene. IARC Monographs on the Evaluation of Carcinogenic Risks to Humans, vol. 82. Lyon, France, IARC, 2002
4. Chuang SC, La Vecchia C, Boffetta P. Liver cancer: descriptive epidemiology and risk factors other than HBV and HCV infection. Cancer Lett 2009; 286: 9-14,
5. Hu G, Tuomilehto J, Pukkala E, Hakulinen T, Antikainen R, Vartiainen E, Jousilahti P. Joint effects of coffee consumption and serum gammaglutamyltransferase on the risk of liver cancer. Hepatology 2008; 48: 129136.
6. Lai GY, Weinstein SJ, Albanes D, Taylor PR, McGlynn KA, Virtamo J, Sinha R, Freedman ND. The association of coffee intake with liver cancer incidence and chronic liver disease mortality in male smokers. Br J Cancer. 2013; 109:1344-51.
7. World Cancer Research Fund / American Institute for Cancer Research. Food, Nutrition, Physical Activity, and the Prevention of Cancer: a Global Perspective. Washington DC, USA, AICR; 2007
8. Bravi F, Bosetti C, Tavani A, Gallus S, La Vecchia C. Coffee Reduces Risk for Hepatocellular Carcinoma: An Updated Meta-Analysis. Clin Gastroenterol Hepatol 2013; 11:1413-21,
9. Fon Sing M, Yang WS, Gao S, Gao J, Xiang YB. Epidemiological studies of the association between tea drinking and primary liver cancer: a metaanalysis. Eur J Cancer Prev 2011; 20: 157-65.
10. Riboli E, Hunt KJ, Slimani N, Ferrari P, Norat T, Fahey M, Charrondière UR, Hémon B, Casagrande C, Vignat J, Overvad K, Tjønneland A, et al. European Prospective Investigation into Cancer and Nutrition (EPIC): study populations and data collection. Public Health Nutr 2002; 5: 1113-1124.
11. Slimani N, Fahey M, Welch AA, Wirfält E, Stripp C, Bergström E, Linseisen J, Schulze MB, Bamia C, Chloptsios Y, Veglia F, Panico S, et al. Diversity of dietary patterns observed in the European Prospective Investigation into Cancer and Nutrition (EPIC) project. Public Health Nutr 2002; 5: 1311-1328.
12. Slimani N, Deharveng G, Unwin I, Southgate DA, Vignat J, Skeie G, Salvini S, Parpinel M, Møller A, Ireland J, Becker W, Farran A, et al. The EPIC nutrient database project (ENDB): a first attempt to standardize nutrient databases across the 10 European countries participating in the EPIC study. Eur J Clin Nutr 2007; 61: 1037-1056.
13. Trichopoulos D, Bamia C, Lagiou P, Fedirko V, Trepo E, Jenab M, Pischon T, Nöthlings U, Overved K, Tjønneland A, Outzen M, Clavel-Chapelon F, et al. Hepatocellular carcinoma risk factors and disease burden in a European cohort: a nested case-control study. J Natl Cancer Inst 2011; 103: 1686-95.
14. Ohfuji S, Fukushima W, Tanaka T, Habu D, Tamori A, Sakaguchi H, Takeda T, Kawada N, Seki S, Nishiguchi S, Shiomi S, Hirota Y. Coffee consumption
and reduced risk of hepatocellular carcinoma among patients with chronic type C liver disease: a case-control study. Hepatol Res 2006; 36: 201-208.
15. Tanaka K, Hara M, Sakamoto T, Higaki Y, Mizuta T, Eguchi Y, Yasutake T, Ozaki I, Yamamoto K, Onohara S, Kawazoe S, Shigematsu H, et al. Inverse association between coffee drinking and the risk of hepatocellular carcinoma: a case-control study in Japan. Cancer Sci 2007; 98: 214-218.
16. Leung WW, Ho SC, Chan HL, Wong V, Yeo W, Mok TS.. Moderate coffee consumption reduces the risk of hepatocellular carcinoma in hepatitis B chronic carriers: a case- control study. J Epidemiol Community Health 2011; 65: 556-558.
17. Inoue M, Yoshimi I, Sobue T, Tsugane S; JPHC Study Group. Influence of coffee drinking on subsequent risk of hepatocellular carcinoma: a prospective study in Japan. J Natl Cancer Inst 2005; 97: 293-300
18. Kurozawa Y, Ogimoto I, Shibata A, Nose T, Yoshimura T, Suzuki H, Sakata R, Fujita Y, Ichikawa S, Iwai N, Tamakoshi A; JACC Study Group. Coffee and risk of death from hepatocellular carcinoma in a large cohort study in Japan. Br J Cancer 2005; 93: 607-610.
19. Shimazu T, Tsubono Y, Kuriyama S, Ohmori K, Koizumi Y, Nishino Y, Shibuya D, Tsuji I. Coffee consumption and the risk of primary liver cancer: pooled analysis of two prospective studies in Japan. Int J Cancer 2005; 116: 150-154
20. Ohishi W, Fujiwara S, Cologne JB, Suzuki G, Akahoshi M, Nishi N, Takahashi I, Chayama K. Risk factors for hepatocellular carcinoma in a Japanese population: a nested case-control study. Cancer Epidemiol Biomarkers Prev 2008; 17: 846-854.
21. Inoue M, Kurahashi N, Iwasaki M, Shimazu T, Tanaka Y, Mizokami M, Tsugane S; Japan Public Health Center-Based Prospective Study Group. Effect of coffee and green tea consumption on the risk of liver cancer: cohort analysis by hepatitis virus infection status. Cancer Epidemiol Biomarkers Prev 2009; 18: 1746-1753
22. Johnson S, Koh WP, Wang R, Govindarajan S, Yu MC, Yuan JM. Coffee consumption and reduced risk of hepatocellular carcinoma: findings from the Singapore Chinese Health Study. Cancer Causes Control 2011; 22: 503-510
23. Gallus S, Bertuzzi M, Tavani A, Bosetti C, Negri E, La Vecchia C, Lagiou P, Trichopoulos D.. Does coffee protect against hepatocellular carcinoma? Br J Cancer 2002; 87: 956-959.
24. Gelatti U, Covolo L, Franceschini M, Pirali F, Tagger A, Ribero ML, Trevisi P, Martelli C, Nardi G, Donato F; Brescia HCC Study Group.. Coffee consumption reduces the risk of hepatocellular carcinoma independently of its aetiology: a case-control study. J Hepatol 2005; 42: 528-534.
25. Montella M, Polesel J, La Vecchia C, Dal Maso L, Crispo A, Crovatto M, Casarin P, Izzo F, Tommasi LG, Talamini R, Franceschi S. Coffee and tea consumption and risk of hepatocellular carcinoma in Italy. Int J Cancer 2007; 120: 1555-1559.
26. Kuper H, Tzonou A, Kaklamani E, Hsieh CC, Lagiou P, Adami HO, Trichopoulos D, Stuver SO. Tobacco smoking, alcohol consumption and their interaction in the causation of hepatocellular carcinoma. Int J Cancer 85: 2000; 498-502.
27. Kanazir M, Boricic I, Delic D, Tepavcevic DK, Knezevic A, Jovanovic T, Pekmezovic T. Risk factors for hepatocellular carcinoma: a case-control study in Belgrade (Serbia). Tumori 2010; 96: 911-7.
28. Wang JY, Zhu L, Wu DL, et al. Matched case-control study on factors for main cancer in a low incidence area of Jiangsu province. Chin J Cancer Prev Treat 2008; 15: 565-568
29. Ui A, Kuriyama S, Kakizaki M, Sone T, Nakaya N, Ohmori-Matsuda K, Hozawa A, Nishino Y, Tsuji I. Green tea consumption and the risk of liver cancer in Japan: the Ohsaki cohort study, Cancer Causes Control 2009; 20: 1939-1945
30. Chen SD, Zhang GQ, Lian JH. A case-control study of primary liver cancer in Guangdong province of China. Chin J Public Health 1999; 18: 231-232.
31. Mu LN, Zhou XF, Ding BG, Wang RH, Zhang ZF, Chen CW, Wei GR, Zhou XM, Jiang QW, Yu SZ. A case-control study on drinking green tea and decreasing risk of cancers in the alimentary canal among cigarette smokers and alcohol drinkers. Chin J Epidemiol 2003; 24: 192-195.
32. Mu LN, Zhou XF, Ding BG, Wang RH, Zhang ZF, Jiang QW, Yu SZ. Study on the protective effect of green tea on gastric, liver and esophageal cancers. Chin J Prev Med 2003; 37: 171-173
33. Nakachi K, Matsuyama S, Miyake S, Suganuma M, Imai K.. Preventive effects of drinking green tea on cancer and cardio vascular disease: epidemiological evidence for multiple targeting prevention. Biofactors., 2000; 13: 49-54.
34. Nagano J, Kono S, Preston DL, Mabuchi K. A prospective study of green tea consumption and cancer incidence, Hiroshima and Nagasaki (Japan). Cancer Causes Control 2001; 12: 501-8.
35. Tang BM, Bian JC, Wang QJ, et al. The case-control study for the risk factors of HCC in Luoyang. Modern Prev Med 2002; 29: 723-725
36. Ruhl CE, Everhart JE. Coffee and tea consumption are associated with a lower incidence of chronic liver disease in the United States. Gastroenterology 2005;129: 1928-36
37. Saab S, Mallam D, Cox GA 2nd, Tong MJ. Impact of coffee on liver diseases: a systematic review. Liver Int. $2014 ; 34: 495-504$. doi: 10.1111/liv.12304. Epub 2013 Sep 15.
38. Corrao G, Zambon A, Bagnardi V, D'Amicis A, Klatsky A Coffee, Caffeine, and the Risk of Liver Cirrhosis Annals of Epidemiology 2001: 11:,458-465
39. Huxley R, Lee CM, Barzi F, Timmermeister L, Czernichow S, Perkovic V, Grobbee DE, Batty D, Woodward M. Coffee. Decaffeinated Coffee, and Tea Consumption in Relation to Incident Type 2 Diabetes Mellitus. A Systematic Review With Meta-analysis. Arch Intern Med 2009; 169: 2053-2063.
40. Torres DM, Harrison SA. Is It Time to Write a Prescription for Coffee?

Coffee and Liver Disease. Gastroenterology 2009; 144: 670-672
41. Cavin C, Holzhaeuser D, Scharf G, Constable A, Huber WW, Schilter B. Cafestol and kahweol, two coffee specific diterpenes with anticarcino-Genic activity. Food Chem Toxicol 2002; 40: 1155-63
42. Moreira DP, Monteiro MC, Ribeiro-Alves M, Donangelo CM, Trugo LC. Contribution of chlorogenic acids to the iron-reducing activity of coffee beverages. J Agric Food Chem 2005; 53: 1399-402
43. Darvesh AS, Bishayee A. Chemopreventive and therapeutic potential of tea polyphenols in hepatocellular cancer. Nutr Cancer 2013; 65:329-44

Table 1. Distribution of study participants by quintiles ${ }^{\text {a }}$ of coffee intake, as well as, by country of origin and baseline characteristics. The European Prospective Investigation into Cancer and nutrition study.

|  | Quintiles of coffee intake (ml/day) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Q1 | Q2 | Q3 | Q4 | Q5 |
| Median (range): males | $\begin{gathered} 17.3 \\ (0 \text { to } 83.3) \end{gathered}$ | $\begin{gathered} 140.5 \\ (83.3 \text { to } 200.4) \end{gathered}$ | $\begin{gathered} 375 \\ (200.5 \text { to } 476.9) \end{gathered}$ | $\begin{gathered} 556.4 \\ (477.2 \text { to } 830.4) \end{gathered}$ | $\begin{gathered} 1000 \\ (831.3 \text { to } 4500) \end{gathered}$ |
| Median (range): females | $3.8$ $\begin{gathered} 5.0 \\ (0 \text { to } 60) \end{gathered}$ | $\begin{gathered} 125.9 \\ (60 \text { to } 191.9) \end{gathered}$ | $290.4$ <br> (191.9 to 375) | $\begin{gathered} 500 \\ (375 \text { to } 580.2) \end{gathered}$ | $856.9$ |
| $\text { Country }{ }^{\mathrm{b}} \text { (\%) }$ |  |  |  |  |  |
| France | 22.1 | 21.9 | 27.0 | 16.9 | 12.0 |
| Italy | 35.6 | 56.7 | 7.4 | 0.3 | 0.04 |
| Spain | 44.7 | 34.0 | 17.8 | 2.8 | 0.7 |
| UK | 26.7 | 20.5 | 10.9 | 21.7 | 20.1 |
| Netherlands | 8.2 | 5.0 | 21.8 | 29.0 | 36.0 |
| Greece | 22.3 | 43.8 | 26.1 | 6.3 | 1.5 |
| Germany | 11.9 | 11.2 | 32.4 | 27.9 | 16.7 |
| Sweden | 7.7 | 10.5 | 27.7 | 32.4 | 21.7 |
| Denmark | 7.4 | 6.6 | 3.1 | 27.0 | 55.9 |
| Norway | 14.4 | 7.2 | 30.8 | 25.9 | 21.7 |
| Age (years): Mean (SD) | 50 (11) | 52 (10) | 51 (9) | 51 (10) | 52 (9) |
| $\mathrm{Sex}^{\text {b }}$ (\%) |  |  |  |  |  |
| males | 20.4 | 19.6 | 21.5 | 18.5 | 20.0 |
| females | 20.1 | 21.2 | 18.7 | 20.4 | 19.6 |
| Self-reported diabetes mellitus ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| No | 20.2 | 21.0 | 19.6 | 19.7 | 19.5 |


| Yes | 24.7 | 24.9 | 19.3 | 15.6 | 15.5 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Unknown | 16.2 | 12.0 | 18.3 | 25.5 | 28.0 |
| Educationallevel ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| None / Primary | 23.2 | 26.9 | 17.9 | 14.6 | 17.4 |
| Technical/professional/secondary | 17.6 | 17.3 | 20.4 | 22.5 | 22.3 |
| University degree/longer | 20.1 | 18.9 | 21.3 | 21.4 | 18.4 |
| Not specified | 26.7 | 22.5 | 12.4 | 21.2 | 17.3 |
| Smoking ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| Never | 24.8 | 23.0 | 19.6 | 19.0 | 13.6 |
| Former | 18.0 | 19.0 | 20.5 | 21.6 | 21.0 |
| Current | 12.6 | 18.0 | 17.9 | 19.5 | 32.0 |
| Unknown | 20.7 | 19.9 | 25.8 | 19.2 | 14.3 |
| Ethanol Intake (g/day) ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| low ${ }^{\text {c }}$ | 23.5 | 21.2 | 19.9 | 18.6 | 16.8 |
| moderate ${ }^{\text {c }}$ | 16.3 | 19.8 | 19.3 | 21.8 | 22.8 |
| high ${ }^{\text {c }}$ | 17.1 | 21.6 | 18.8 | 19.4 | 23.2 |
| Body Mass Index ${ }^{\text {b }}$ (kg/m ${ }^{2}$ ) (\%) |  |  |  |  |  |
| ce25 | 20.1 | 19.4 | 20.0 | 21.0 | 19.5 |
| $>25-<30$ | 19.2 | 21.3 | 19.5 | 19.3 | 20.6 |
| $>=30$ | 22.8 | 24.5 | 18.0 | 16.3 | 18.4 |
| Physical Activity Index ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| moderately inactive ${ }^{\text {d }}$ | 22.5 | 24.6 | 18.8 | 17.6 | 16.5 |
| moderately active ${ }^{\text {d }}$ | 18.4 | 18.6 | 18.5 | 21.3 | 23.2 |
| unknown | 14.2 | 7.5 | 28.7 | 26.3 | 23.3 |
| Energy Intake (kcal/day) : mean (SD) | 2051.4 (686.0) | 2109.6 (677.0) | 2062.7 (661.1) | 2069.8 (662.9) | $2130.2(702.3)$ |


| Total(Number) | 98,148 | 100,953 | 95,231 | 96,413 | 96,054 |
| :--- | :--- | :--- | :--- | :--- | :--- |

${ }^{\text {a }}$ Sex-specific quintiles in the overall cohort
${ }^{\mathrm{b}}$ Numbers in cells do not always add up to $100 \%$ horizontally due to rounding
${ }^{c}$ Ethanol Intake: Low: Males: $<10 \mathrm{~g} /$ day / Females: $<5 \mathrm{~g} /$ day; Moderate: Males:10g/day to $<40 \mathrm{~g} / \mathrm{day} /$ Females:5g/day to 20g/day; High: Males: $>40 \mathrm{~g} /$ day; Females $:>20 \mathrm{~g} /$ day.
${ }^{d}$ Based on occupational physical activity and physical activity during cycling and sports

Table 2. Distribution of study participants by categories ${ }^{\text {a }}$ of tea intake, as well as, by country of origin and baseline characteristics. The European Prospective Investigation into Cancer and nutrition study.

|  | Categories of tea intake (ml/day) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | C1 | C2 | C3 | C4 | C5 |
| Median (range): males | 0 | $\begin{gathered} 3.3 \\ (0.3 \text { to 16.4) } \end{gathered}$ | $\begin{gathered} 37 \\ \text { (16.4 to 106.8) } \end{gathered}$ | $\begin{gathered} 250 \\ (107.1 \text { to 475) } \end{gathered}$ | $\begin{gathered} 855 \\ (480 \text { to } 2375) \end{gathered}$ |
| Median (range): females | 0 | $\begin{gathered} 8.2 \\ (0.0006 \text { to } 28.6) \end{gathered}$ | $\begin{gathered} 87.9 \\ (28.6 \text { to 190) } \end{gathered}$ | $\begin{gathered} 360 \\ (190.4 \text { to 475) } \end{gathered}$ | $\begin{gathered} 855 \\ (475.1 t 05937.5) \end{gathered}$ |
| Country ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| France | 41.0 | 10.2 | 16.9 | 19.0 | 13.0 |
| Italy | 45.1 | 23.8 | 24.8 | 5.9 | 0.4 |
| Spain | 96.0 | 0.5 | 1.9 | 1.5 | 0.1 |
| UK | 2.8 | 14.3 | 11.0 | 34.4 | 37.6 |
| Netherlands | 10.5 | 8.0 | 23.1 | 44.3 | 14.2 |
| Greece | 46.2 | 44.5 | 8.0 | 1.2 | 0.1 |
| Germany | 24.1 | 6.0 | 35.4 | 26.3 | 8.2 |
| Sweden | 52.4 | 16.6 | 13.3 | 12.6 | 5.1 |
| Denmark | 16.1 | 19.3 | 20.8 | 10.0 | 33.8 |
| Norway | NA | NA | NA | NA | NA |
| Age a(years): Mean (SD) | 53 (9) | 50 (11) | 50 (10) | 51 (11) | 53 (11) |
| $\mathrm{Sex}^{\text {b }}$ (\%) |  |  |  |  |  |
| males | 35.3 | 17.4 | 15.2 | 18.2 | 14.0 |
| females | 33.3 | 16.7 | 16.8 | 17.5 | 15.7 |
| Self reported diabetes mellitus ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| No | 33.6 | 17.1 | 16.5 | 17.9 | 15.0 |


| Yes | 47.5 | 15.4 | 13.7 | 12.9 | 10.5 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Unknown | 31.2 | 13.9 | 12.6 | 17.5 | 24.9 |
| Educationallevel ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| None / Primary | 53.1 | 16.5 | 13.4 | 8.8 | 8.1 |
| Technical/professional/secondary | 27.7 | 17.6 | 18.1 | 20.3 | 16.4 |
| University degree/longer | 24.1 | 17.0 | 17.7 | 22.6 | 18.7 |
| Not specified | 11.2 | 12.9 | 11.2 | 30.2 | 34.5 |
| Smoking ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| Never | 33.3 | 16.6 | 16.4 | 18.5 | 15.2 |
| Former | 29.8 | 15.7 | 16.5 | 20.1 | 17.8 |
| Current | 40.8 | 19.2 | 15.8 | 12.9 | 11.3 |
| Unknown | 29.9 | 16.8 | 15.1 | 20.0 | 18.2 |
| Ethanol Intake (g/day $)^{\text {b }}$ (\%) |  |  |  |  |  |
| low ${ }^{\text {c }}$ | 35.1 | 17.8 | 15.5 | 17.3 | 14.3 |
| moderate ${ }^{\text {c }}$ | 30.8 | 16.3 | 17.2 | 19.0 | 16.7 |
| high $^{\text {c }}$ | 37.8 | 15.5 | 16.8 | 16.0 | 13.9 |
| Body Mass Index (kg/m $\left.{ }^{2}\right)^{\mathrm{b}}(\%)$ |  |  |  |  |  |
| < $=25$ | 28.1 | 16.7 | 17.1 | 20.5 | 17.6 |
| $>25-<30$ | 37.5 | 17.1 | 15.8 | 16.0 | 13.6 |
| >=30 | 46.2 | 17.5 | 14.5 | 11.9 | 10.0 |
| Physical Activity Index ${ }^{\text {b }}$ (\%) |  |  |  |  |  |
| moderately inactive ${ }^{\text {d }}$ | 38.6 | 17.1 | 15.1 | 15.6 | 13.7 |
| moderately active ${ }^{\text {d }}$ | 28.6 | 17.1 | 17.8 | 19.8 | 16.8 |
| unknown | 12.9 | 9.3 | 19.8 | 38.3 | 19.7 |
| Energy Intake (kcal/day) : mean (SD) | 2154.6 (691.5) | 2053.1(682.8) | 2140.3 (698.3) | 2084.2 (651.5) | 2132.3 (672.2) |


| Total $^{\mathrm{e}}$ (Number) | 153,097 | 76,393 | 73,424 | 79,872 | 68,135 |
| :--- | :--- | :--- | :--- | :--- | :--- |

${ }^{\text {a }}$ Categories of tea intake were defined as: non-consumers ( C 1 ); and the sex-specific quartiles among tea consumers (C2-C5) in the overall cohort.
${ }^{b}$ Numbers in cells do not always add up to $100 \%$ horizontally due to rounding
${ }^{c}$ Ethanol Intake: Low: Males: $<10 \mathrm{~g} /$ day / Females: $<5 \mathrm{~g} /$ day; Moderate: Males: $10 \mathrm{~g} /$ day to $<40 \mathrm{~g} /$ day / Females: $5 \mathrm{~g} /$ day to $20 \mathrm{~g} /$ day; High: Males: $>40 \mathrm{~g} /$ day; Females: $>20 \mathrm{~g} /$ day.
${ }^{d}$ Based on occupational physical activity and physical activity during cycling and sports
${ }^{\mathrm{e}}$ Norway is excluded due to lack of information about tea intake

Table 3. Hazard Ratios (HR) and 95\% Confidence Intervals (CI) of hepatocellular carcinoma (HCC) incidence according to quintiles ${ }^{\text {a }}$ of coffee and categories ${ }^{\mathbf{b}}$ of tea intake in the European Prospective Investigation into Cancer and nutrition cohort.

|  | $\begin{gathered} \mathrm{HCC} \\ \text { cases/N} \end{gathered}$ | Model $1^{\text {c }}$ |  | Model $2{ }^{\text {d }}$ |  | Model $3^{\text {e }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | HR | 95\% CL | HR | 95\% CI | HR | 95\% CL |
| Coffee Intake (quintiles) |  |  |  |  |  |  |  |
| Q1 | 47/98,148 | 1 |  | 1 |  | 1 |  |
| Q2 | 49/100,953 | 0.96 | (0.63 to 1.44) | 0.91 | (0.60 to 1.38$)$ | 0.85 | (0.56 to 1.29) |
| Q3 | 38/95,231 | 0.75 | (0.46 to 1.19) | 0.71 | (0.44 to 1.14) | 0.63 | (0.39 to 1.02) |
| Q4 | 36/96,413 | 0.67 | (0.41 to 1.10$)$ | 0.61 | (0.37 to 1.01 ) | 0.49 | (0.29 to 0.82) |
| Q5 | 31/96,054 | 0.48 | (0.28 to 0.82) | 0.38 | ( 0.22 to 0.66 ) | 0.28 | ( 0.16 to 0.50 ) |
| $P$-trend |  | 0.004 |  | $<0.001$ |  | $<0.001$ |  |
| Tea Intake (categories) |  |  |  |  |  |  |  |
| C1 | 85/153,097 | 1 |  | 1 |  | 1 |  |
| C 2 | 37/76,393 | 0.96 | (0.62 to 1.49) | 1.04 | (0.67 to 1.61$)$ | 1.05 | (0.68 to lo 1.63 ) |
| C3 | 35/73,424 | 0.90 | ( 0.58 to 1.41 ) | 1.01 | ( 0.65 to 1.58 ) | 0.98 | ( 0.63 to 1.53 ) |
| C4 | 24/79,872 | 0.68 | (0.40 to 1.16) | 0.82 | (0.48 to 1.41) | 0.71 | (0.41 to 1.23) |
| C5 | 18/68,135 | 0.49 | (0.27 to 0.89) | 0.60 | (0.32 to 1.10 ) | 0.41 | (0.22 to 0.78) |
| $P$-trend |  | 0.01 |  | 0.06 |  | 0.003 |  |

Abbreviations: HR, Hazard ratio; CI, Confidence Interval; HCC, Hepatocellular carcinoma
${ }^{\text {a }}$ Sex-specific quintiles (Q1-Q5) in the overall cohort
${ }^{\mathrm{b}}$ Categories of tea intake were defined as: non-consumers (C1); and the sex-specific quartiles among tea consumers (C2-C5) in the overall cohort.
${ }^{\text {c }}$ Stratified for age at recruitment (in 1-year intervals) and for centre. Adjusted for sex. Norway was not included in analyses concerning tea intake as this information was not available.
${ }^{\mathrm{d}}$ Stratified for age at recruitment (in 1-year intervals) and for centre. Adjusted for sex, diabetes mellitus (self reported at enrolment), education, body mass index, smoking, physical activity, alcohol intake, categorically as in Table 2 and energy intake (ordered in quintiles). Norway was not included in analyses concerning tea intake as this information was not available.
${ }^{\mathrm{e}}$ As model 2 but simultaneously including tea and coffee intake. Norway was not included in analyses concerning tea intake as this information was not available.

