Imaging markers of disability in aquaporin-4 immunoglobulin G seropositive neuromyelitis optica: a graph theory study

Claudia Chien,1,2 Frederike Cosima Oertel,1,2 Nadja Siebert,1,2,3 Hanna Zimmermann,1,2 Susanna Asseyer,1,2 Joseph Kuchling,1,2,3 Michael Scheel,1,2,4 Klemens Ruprecht,3 Judith Bellmann-Strobl,1,2,3 Friedemann Paul1,2,3,* and Alexander U. Brandt1,2,5,*

These authors contributed equally to this work.

Neuromyelitis optica spectrum disorders lack imaging biomarkers associated with disease course and supporting prognosis. This complex and heterogeneous set of disorders affects many regions of the central nervous system, including the spinal cord and visual pathway. Here, we use graph theory-based multimodal network analysis to investigate hypothesis-free mixed networks and associations between clinical disease with neuroimaging markers in 40 aquaporin-4-immunoglobulin G antibody seropositive patients (age \(\mu = 48.16 \pm 14.3\) years, female:male = 36:4) and 31 healthy controls (age \(\mu = 45.92 \pm 13.3\) years, female:male = 24:7). Magnetic resonance imaging measures included total brain and deep grey matter volumes, cortical thickness and spinal cord atrophy. Optical coherence tomography measures of the retina and clinical measures comprised of clinical attack types and expanded disability status scale were also utilized. For multimodal network analysis, all measures were introduced as nodes and tested for directed connectivity from clinical attack types and disease duration to systematic imaging and clinical disability measures. Analysis of variance, with group interactions, gave weights and significance for each nodal association (hyperedges). Connectivity matrices from 80% and 95% F-distribution networks were analyzed and revealed the number of combined attack types and disease duration as the most connected nodes, directly affecting changes in several regions of the central nervous system. Subsequent multivariable regression models, including interaction effects with clinical parameters, identified associations between decreased nucleus accumbens (\(\beta = -0.85, P = 0.021\)) and caudate nucleus (\(\beta = -0.61, P = 0.011\)) volumes with higher combined attack type count and longer disease duration, respectively. We also confirmed previously reported associations between spinal cord atrophy with increased number of clinical myelitis attacks. Age was the most important factor associated with normalized brain volume, pallidum volume, cortical thickness and the expanded disability status scale score. The identified imaging biomarker candidates warrant further investigation in larger-scale studies. Graph theory-based multimodal networks allow for connectivity and interaction analysis, where this method may be applied in other complex heterogeneous disease investigations with different outcome measures.

1 Experimental and Clinical Research Center, Max Delbrück Center for Molecular Medicine and Charité –Universitätsmedizin Berlin, Corporate Member of Freie Universität Berlin, Humboldt-Universität zu Berlin, and Berlin Institute of Health, Germany
2 NeuroCure Clinical Research Center, Charité –Universitätsmedizin Berlin, Corporate Member of Freie Universität Berlin, Humboldt-Universität zu Berlin, and Berlin Institute of Health, Germany
3 Department of Neurology, Charité –Universitätsmedizin Berlin, Corporate Member of Freie Universität Berlin, Humboldt-Universität zu Berlin, and Berlin Institute of Health, Germany
4 Department of Neuroradiology, Charité –Universitätsmedizin Berlin, Corporate Member of Freie Universität Berlin, Humboldt-Universität zu Berlin, and Berlin Institute of Health, Germany
5 Department of Neurology, University of California, Irvine, CA, USA

Received May 21, 2019. Revised August 14, 2019. Accepted October 1, 2019. Advance Access publication October 16, 2019
© The Author(s) (2019). Published by Oxford University Press on behalf of the Guarantors of Brain.
This is an Open Access article distributed under the terms of the Creative Commons Attribution Non-Commercial License (http://creativecommons.org/licenses/by-nc/4.0/), which permits non-commercial re-use, distribution, and reproduction in any medium, provided the original work is properly cited. For commercial re-use, please contact journals.permissions@oup.com
Introduction

Neuromyelitis optica spectrum disorders (NMOSD) are inflammatory diseases of the central nervous system defined by a clinical spectrum of optic neuritis (ON), myelitis, and more rarely, brainstem and cerebral attacks (Wingerchuk et al., 2015). The majority of patients (about 70%) are seropositive for immunoglobulin (Ig)G antibodies against the astrocytic water channel aquaporin-4 (AQP4-IgG+) (Mori et al., 2018; Cook et al., 2019). Magnetic resonance imaging (MRI) identification of T2-hyperintense brain lesions is useful for initial differential diagnosis and during acute attacks, but brain lesions are often non-specific for disability or prognosis in NMOSD (Kremer et al., 2015).

Several imaging markers have been proposed for quantifying direct attack-related damage, including spinal cord atrophy after myelitis (Chien et al., 2018b), neuro-axonal damage in the retina assessed by optical coherence tomography (OCT) after ON (Schmidt et al., 2017; Oberwahrenbrock et al., 2018; Oertel et al., 2018) and T2-hyperintense lesions in symptomatic brainstem syndromes (Kim et al., 2015).

However, it is unclear if the NMOSD disease course also leads to covert or diffuse and clinically relevant central nervous system damage, which can be systematically assessed by imaging biomarkers, for example, in the grey matter (Pache et al., 2016) or the retina (Oertel et al., 2018). Normal MRI volumetric deep grey matter (DGM)
measures have been reported by some (Finke et al., 2016), and it has been difficult to attribute clinical disability in NMOSD to systematic measures (Mealy et al., 2019). Others reported cortical grey matter atrophy (Kim et al., 2016) or DGM affection with clinical relevance, i.e. for cognition (Kim et al., 2017; Oertel et al., 2019).

Complex dependencies of clinical observations and central nervous system measures can be captured using graph theory models relying on a network of interactions rather than individual pairwise analyses (Lambiotte et al., 2019). Graph theory models distinct anatomical brain regions as nodes with connections represented by edges, which can be exploited to calculate the strength of interaction between multiple regions (Fleischer et al., 2017). Graph theory-based network analysis is typically applied in a single modality network, e.g. magnetic resonance imaging measures. Here, we extend this approach to evaluate multimodal connectivity (Heath and Sioson, 2009b) between distinct systematic MRI, OCT and clinical measures to identify imaging correlates of disease activity in AQP4-IgG+ NMOSD patients.

Materials and methods

Participants

Cross-sectional data from 65 participants from an ongoing observational study at the Experimental and Clinical Research Center and the NeuroCure Clinical Research Center, Charité-Universitätsmedizin Berlin, were screened for inclusion in this retrospective study. Inclusion criteria were a minimum age of 18 years, a diagnosis of NMOSD according to the 2015 International Consensus Diagnostic Criteria (Wingerchuk et al., 2015) with seropositivity for AQP4-IgG in a cell-based assay (Euroimmune, Lübeck, Germany) at any time during the course of the disease, and availability of MRI, OCT images and clinical data. No restrictions were made as to type of attack preventing therapy. Ophthalmologic comorbidities (e.g. glaucoma, myopia >5dpt) were excluded in OCT analysis only. We included data from 40 patients with NMOSD and 31 age- and sex-matched healthy controls (HC) from our research database (Table 1). Disease duration was calculated from the first clinical symptom to the date of the MRI examination. Clinical assessments consisted of the expanded disability status scale (EDSS), number of ON attacks, myelitis attacks, other NMOSD-related clinical attacks (i.e. area postrema or brainstem syndrome) and combined attacks (i.e. simultaneous myelitis and ON or myelitis and brainstem syndrome). The combined attacks were counted also as separate types of attacks, for example, a patient with a simultaneous myelitis and brainstem attack would have a count of one myelitis and one other attack; however, the total number of attacks would remain as one for that patient. The study was approved by the ethics committee at the Charité-Universitätsmedizin Berlin, Germany (EA1/131/09, EA1/163/12 and EA1/077/11) and was conducted in accordance to the Declaration of Helsinki in its current

Table 1 Participant demographics

<table>
<thead>
<tr>
<th></th>
<th>HC</th>
<th>AQP4-IgG+ NMOSD</th>
<th>Test statistics</th>
</tr>
</thead>
<tbody>
<tr>
<td>Number of subjects [n] (total = 71)</td>
<td>31</td>
<td>40</td>
<td>NA</td>
</tr>
<tr>
<td>Age [years] (mean ± SD)</td>
<td>45.92 ± 13.3</td>
<td>48.16 ± 14.3</td>
<td>$\chi^2 = 0.7164$, $P = 0.3973$</td>
</tr>
<tr>
<td>Female:male [n] (%female)</td>
<td>24:7 (77%)</td>
<td>36.4 (92%)</td>
<td>$\chi^2 = 1.2597$, $P = 0.2617$</td>
</tr>
<tr>
<td>Disease duration [years] [mean (range)]</td>
<td>NA</td>
<td>4.6 (0.47–28.06)</td>
<td>NA</td>
</tr>
<tr>
<td>Attack types [n] [median (range)]</td>
<td>NA</td>
<td>ON: 1 (0–14) Myelitis: 1 (0–11) Other: 0 (0–2) Combined: 0 (0–2) Total: 3 (1–14)</td>
<td>NA</td>
</tr>
<tr>
<td>EDSS [median (range)]</td>
<td>NA</td>
<td>4 (0–6.5)</td>
<td>NA</td>
</tr>
<tr>
<td>Attack preventing therapies [n] (%)</td>
<td>NA</td>
<td>AZA: 8 (20) BLM: 1 (2.5) GALA: 1 (2.5) MTX: 1 (2.5) MMF: 1 (2.5) RIX: 17 (42.5) N: 11 (27.5)</td>
<td>NA</td>
</tr>
</tbody>
</table>

Clinical demographics of study cohort, where attack preventing therapy counts for all patients were included prior to the MRI date. AQP4-IgG+ NMOSD = aquaporin-4 immunoglobulin-G seropositive neuromyelitis optica spectrum disorders; AZA = azathioprine; BLM = belimumab; EDSS = expanded disability status scale; GALA = glatiramer acetate; HC = healthy control; MMF = mycophenolate mofetil; MTX = methotrexate; N = none/unknown; ON = optic neuritis; RIX = rituximab.
applicable version. All participants provided written informed consent.

**Magnetic resonance imaging**

All MRI scans were performed on a 3-T Siemens MAGNETOM Trio Tim (Erlangen, Germany) scanner at the Berlin Center for Advanced Neuroimaging. The MRI protocol for this study included: (i) a T1-weighted 3D magnetization-prepared rapid gradient echo brain MRI (1 × 1 × 1 mm resolution, repetition time (TR) = 1900 ms, echo time (TE) = 3.03 ms), including the upper cervical cord; and (ii) a T2-weighted 3D fluid-attenuated inversion recovery brain MRI (1 × 1 × 1 mm resolution, TR = 6000 ms, TE = 388 ms). All mean upper cervical cord area (MUCCA) measurements were performed at the C2/C3 intervertebral space level using JIM7.0 (Xinapse Systems Ltd, UK) software as described in a previous publication (Chien et al., 2018a). Brain lesion segmentation was performed using the automated lesion segmentation toolbox on fluid-attenuated inversion recovery images (Schmidt et al., 2012) and manually corrected by experienced radiology technicians at our centre. Normalized brain volume (NBV), DGM and brainstem volumes and cortical thickness were obtained from lesion-filled magnetization-prepared rapid gradient echo scans using FSL SIENAX (Smith et al., 2002), FSL FIRST (Patenaude et al., 2011) and the CAT12 toolbox (Dahnke et al., 2013) for MATLAB SPM12. NBV and the total volume of DGM and brainstem structures were normalized by multiplication with the V-Scaling factor, an approximation of head-size given by FSL SIENAX.

**Optical coherence tomography**

Retinal examinations were performed using a Heidelberg Engineering Spectralis SD-OCT (Heidelberg Engineering, Heidelberg, Germany) with automatic real-time (ART) function. Scan quality was checked using the OSCAR-IB Criteria (Schippling et al., 2015) with the Advised Protocol for OCT Study Terminology and Elements APOSTEL recommendations (Cruz-Herranz et al., 2016). 3.4 mm ring scans around the optic nerve head were used to measure the peripapillary retinal nerve fibre layer (12°, 1336 A-scans 16 automated real time tracking (ART) ≤ 100). The ganglion cell inner plexiform (GCIP) layer volume was measured using a 6 mm diameter cylinder around the fovea from macular volume scans (25° × 30°, 61 vertical B-scans, 768 A-scans per B-scan, ART = 15). Layer segmentation was performed using Eye Explorer 1.9.10.0, viewing module 6.3.4.0 (Heidelberg Engineering, Heidelberg, Germany).

**Multimodal network analysis**

All systematic MRI, OCT and clinical measures were considered nodes in a network to test for strength of connection between the number of clinical attacks and disease duration with EDSS, brain, spinal cord and retinal measures. Directed correlational analyses between clinical attack counts and disease duration with all other nodes were performed using ANOVA with Type III sum of squares analyses, adjusted for group, thus accounting for only diseased associations of imaging measures with clinical measures. This type of multimodal network approach was proposed by Heath and Sioson (Heath and Sioson, 2009a, b), where multiple networks can be combined with interactions to represent complex biological processes. We calculated the union of three different networks (MRI measures, OCT measures and clinical measures) to give a subnetwork with ‘hyperedges’ derived from the weights calculated from the directed ANOVA F-statistics (Fig. 1A). A threshold F-statistic value of 1.65 or greater was used (giving a probability of 80% observed difference with a numerator degree of freedom of 2 and denominator degree of freedom of 60) to define hyperedges between nodes in the network. Significant network associations (95% F-distribution) were then investigated by further thresholding the P-value from each ANOVA to be <0.05. Connectivity matrices for each network were constructed using the node pairs with F-statistics and P-values within the thresholds set for 80% and 95% probability. Both networks based on an 80% F-distribution and 95% F-distribution were evaluated for connectivity, which depended on the sum of the degree, output weight, closeness to nodes, and betweenness of nodes, as well as the distance between nodes (Rubinov and Sporns, 2010). These particular graph theory measures were chosen since we employed a multimodal network analysis, different than simple conventional directed network analysis, thus we endeavoured to evaluate the component measures which can lead to analysis of centrality, network motifs and network comparison. Shortest distances between nodes are dependent on and calculated using the edge list with corresponding weights (Opsahl et al., 2010), which are related to betweenness in the nodes. Thus, the higher the sum value for connectivity, the more connected the node is with others in the network, suggesting stronger directed associations of that node with clinical measures. The two nodes showing the highest sum value for connectivity were deemed as the most connected nodes in these multimodal networks. Shortest distances were analysed in a matrix format to visualize the closeness of each type of attack to each imaging and clinical measure. This enabled a straightforward and interpretable post hoc analysis with covariates based on these matrices.

**Statistical analysis**

Proportional group differences of patient demographics in age and sex were calculated using a Chi-squared test. Group differences between NMOSD and HC for all MRI and OCT measures were tested using a Mann–Whitney U-test, without correction for multiple testing. Network connections of MRI, OCT and EDSS measures within a 95% F-distribution that showed close distances (shortest
weighted path length) with attack types and disease duration in the multimodal network analysis were further investigated for effect sizes and changes in the brain, spinal cord and retina of AQP4-IgG+ NMOSD patients, using multivariable linear modelling corrected for by age. Sex was not corrected for, due to the low number of males in the cohort (4 out of 40). For connections with multiple close distances with attack types and/or disease duration, the attack types and/or disease duration were used as interacting covariates. Leave-one-out-cross-validation was used to evaluate prediction error of multivariable linear models, where low overall mean squared error (mean squared error < 20% of mean) over several fold tests would indicate low prediction error. All statistical analyses and plots were produced using R version 3.4.0. Statistical significance in all tests was set to a P-value < 0.05, reflecting true test statistics (confidence intervals) and correlation coefficients (effect sizes).

Data availability

All data from this study can be shared at the request of other investigators for purposes of replicating procedures and results.

Results

Group differences in magnetic resonance imaging and optical coherence tomography measures

Participant clinical demographics are given in Table 1. Patient imaging measures were lower than HC (in order of decreasing effect size) for MUCCA, GCIP, pallidum volume, NBV and caudate nucleus volume in Mann–Whitney U group comparisons (Table 2), but higher than HC for T2-brain lesion count and volume. This is in line with an earlier study using in part, overlapping data from the same cohort study (HC: n = 12, NMOSD: n = 32), showing none or minimal volumetric changes in the DGM (Finke et al., 2016).

Multimodal network analysis with 80% probability of observed correlation

We performed a multimodal network analysis with an 80% probability (F-distribution) threshold to minimize false negatives. Analysis of nodal correlations within this network revealed that all MRI, OCT and clinical measures were associated with different attack types and disease duration in AQP4-IgG+ NMOSD (Fig. 1B).

The number of combined attacks and disease duration was shown to be the most connected nodes to NMOSD MRI, OCT and clinical disability (EDSS) nodes. This was calculated using the sum of the degree (number of connections), output (directed weights of hyperedges between nodes), closeness (strength of connections) and correlation coefficients (effect sizes).
betweenness (strength of interaction between two nodes) shown in Table 3.

**Multimodal network analysis with 95% probability of observed correlation**

Next, we increased the probability threshold to 95% to limit false positives. When looking at significant connections between attacks and disease duration with the other MRI, OCT and clinical nodes, we found that the number of combined attacks and disease duration remained the most connected to changes in other measures, shown in Table 4. Meanwhile, disease-related attacks were not found to be linked to brain lesion counts or volume, nor thalamic or hippocampal volumes. The number of other attacks (area postrema or brainstem syndrome) also did not show a significant influence on
Nodes that were removed from the network due to non-statistical significance were the inner nuclear layer, hippocampus, thalamus, putamen, amygdala, brainstem, brain lesion count and volume and the number of other attacks (Fig. 1C).

Multivariable linear analysis

We then used weighted path distances to select the closest connections between imaging and clinical markers (Fig. 2A and B). All connections in the 95% probability network with a path length $< 2$ were tested in multivariable linear regression models, corrected for by age, to derive effect sizes of the associations. Path lengths $< 2$ in a binary shortest distance calculation indicate a direct connection. While for instance, if two nodes are not connected, but share a connection with another node, the shortest distance (path length) between these two nodes would be two (Opsahl et al., 2010). Thus, we specifically chose directly connected nodes for further investigation. Here, NBV, pallidum volume, cortical thickness and EDSS score were only significantly associated with patient age, but not with any clinical measures (Table 5).

Meanwhile, the number of combined attacks (i.e. simultaneous myelitis and ON or myelitis and brainstem syndrome) was significantly associated with a decrease in the nucleus accumbens volume ($\beta = -0.61$, $P = 0.011$, Fig. 3B). Using leave-one-out-cross-validation (Sammut and Webb, 2010), it was found over 6-folds that there were prediction errors of 0.0465 (mean square error = 4.2% of the mean) and 1.37 (mean squared error = 15.3% of the mean) for the nucleus accumbens and caudate nucleus multivariable linear models, respectively. Attack type count distributions for this patient cohort are shown in Fig. 3C.

Discussion

Our study identified the following imaging biomarker candidates in AQP4-IgG+ NMOSD patients, where damage seemed to occur indirectly from the attack site:

a. decreases in the nucleus accumbens volume is associated with increased number of combined attacks, and
b. decreases in the caudate nucleus volume associates with increased disease duration.

Regarding attack-site related damage, our study confirmed previously reported associations of myelitis and MUCCA (Chien et al., 2018b) in a 95% probability network, as well as ON and GCIP and peripapillary retinal nerve fibre layer (Oertel et al., 2017) when looking broader in the 80% probability network. Age was the...
most important factor in measures of the NBV, pallidum volume, cortical thickness and EDSS score.

The most intriguing finding of our study is the potential evidence of neurodegeneration in non-attack localized regions in NMOSD, the nucleus accumbens and the caudate nucleus. Previous studies found possible covert damage in the optic radiation and retina, but no occult damage in lesion-free brain regions (Pache et al., 2016; Oertel et al., 2017; Pasquier et al., 2019). Interestingly, an earlier study looked at markers for cognitive impairment in NMOSD patients and found that nucleus accumbens volume was associated with overall cognition and cognitive impairment in these patients. Moreover, cognitively impaired patients with NMOSD showed atrophy in the caudate nucleus (Kim et al., 2017). Although we did not specifically consider cognitive impairment in this study, Kim et al.’s findings alongside ours, serve as independent verification that the nucleus accumbens and caudate nucleus are affected in AQP4-IgG+ NMOSD patients. The nucleus accumbens has been implicated in neuropathic pain and neuropsychiatric conditions and is suggested to be modulated in an aquaporin-4-dependent manner (Wu et al., 2018). Previously, we also found moderate to severely depressed NMOSD patients are more susceptible to insufficiently mediated neuropathic pain and fatigue (Chavarro et al., 2016; Asseyer et al., 2018). Together, these data suggest that nucleus accumbens and caudate nucleus affection in NMOSD may be a true positive finding, warranting further confirmation, although the underlying mechanism remains elusive. Due to these interesting findings, we are encouraged to investigate the potential changes in cognitive, executive and mood functions of NMOSD patients and their associations with DGM volumes calculated from MRI scans.

The main suggested mechanism of damage by aquaporin-4 antibodies in NMOSD is immune-mediated...
Table 5 Multivariable linear models for changes in AQP4-IgG+ NMOSD MRI, OCT and clinical measures

<table>
<thead>
<tr>
<th>Attack Type</th>
<th>ON Attacks</th>
<th>Myelitis Attacks</th>
<th>Other Attacks</th>
<th>Combined Attacks</th>
<th>Total Number of Attacks</th>
<th>Disease Duration (DD)</th>
<th>Interactions</th>
<th>Age</th>
<th>F-statistic</th>
<th>Adjusted $R^2$</th>
<th>Model P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>NBV</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = -0.40, P = 0.159$</td>
<td>$\beta = 0.25, P = 0.247$</td>
<td>Combined: DD</td>
<td>$\beta = -0.67, P = 1.7e-5$</td>
<td>9.922</td>
<td>0.5352</td>
</tr>
<tr>
<td>Caudate nucleus</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = -0.55, P = 0.067$</td>
<td>$\beta = -0.61, P = 0.011$</td>
<td>Combined: DD</td>
<td>$\beta = -0.51, P = 8.0e-4$</td>
<td>8.277</td>
<td>0.4843</td>
</tr>
<tr>
<td>Pallidum</td>
<td>$\beta = -0.15, P = 0.365$</td>
<td>$\beta = 0.03, P = 0.870$</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>Myelitis: ON</td>
<td>$\beta = -0.38, P = 0.013$</td>
<td>3.310</td>
<td>0.1956</td>
<td>0.0215</td>
</tr>
<tr>
<td>Nucleus accumbens</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = -0.85, P = 0.021$</td>
<td>NA</td>
<td>$\beta = -0.43, P = 0.124$</td>
<td>Combined: DD</td>
<td>$\beta = 0.35, P = 0.109$</td>
<td>3.579</td>
<td>0.2497</td>
<td>0.0182</td>
</tr>
<tr>
<td>Cortical thickness</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = 0.18, P = 0.156$</td>
<td>$\beta = 0.042$</td>
<td>$\beta = -0.11, P = 0.0118$</td>
<td>13.88</td>
<td>0.4041</td>
<td>3.4e-5</td>
</tr>
<tr>
<td>MUCCA</td>
<td>$\beta = -4.1, P = 0.023$</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = 0.78, P = 0.511$</td>
<td>$\beta = -0.88, P = 0.348$</td>
<td>$\beta = 0.15, P = 0.907$</td>
<td>Myelitis: ON</td>
<td>$\beta = -0.12, P = 0.449$</td>
<td>2.935</td>
<td>0.4479</td>
<td>0.0180</td>
</tr>
<tr>
<td>pRNFL</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = -0.34, P = 0.107$</td>
<td>$\beta = 0.017, P = 0.415$</td>
<td>$\beta = 0.017, P = 0.1997$</td>
<td>1.748</td>
<td>0.06365</td>
<td>0.1997</td>
</tr>
<tr>
<td>GCIP</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = -0.35, P = 0.105$</td>
<td>$\beta = 0.22, P = 0.294$</td>
<td>$\beta = 0.22, P = 0.294$</td>
<td>2.055</td>
<td>0.08755</td>
<td>0.1542</td>
</tr>
<tr>
<td>EDSS</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>NA</td>
<td>$\beta = 0.25, P = 0.121$</td>
<td>$\beta = 0.50, P = 0.003$</td>
<td>$\beta = 0.50, P = 0.003$</td>
<td>7.252</td>
<td>0.3013</td>
<td>0.0030</td>
</tr>
</tbody>
</table>

Multivariable linear regression models of associations between attack types and disease duration with MRI, OCT and EDSS score, with adjustment for age. Bolded test statistics indicate statistical significance. DD = disease duration; EDSS = expanded disability status scale; GCIP = ganglion cell inner plexiform; MUCCA = mean upper cervical cord volume; NBV = normalized brain volume; ON = optic neuritis; pRNFL = peripapillary retinal nerve fibre layer; $\beta$ = standardized coefficients in multivariable linear regressions.
Figure 3. Associations and distributions of disease-related attacks. Multivariable linear regression models of (A) the nucleus accumbens volume versus the number of combined attacks and (B) the caudate nucleus volume versus the disease duration in AQP4-IgG⁺ NMOSD patients, corrected for age. (C) Distribution of attack type counts in this cohort. Abbreviations: ON = optic neuritis.
attacks against astrocytes, subsequent metabolic breakdown at the attack site and resulting axonal transection throughout structural pathways, such as the anterior visual pathway (Kuchling et al., 2018). In line with attack-related damage, associations of decreased MUCCA with an increased number of clinical myelitis attacks were previously found (Chien et al., 2018b). Regarding retinal measures, previous studies have established peripapillary retinal nerve fibre layer and GCIP as potential measures of NMOSD attack-related damage (Oertel et al., 2018). Our results are in accordance with previous findings, thus strengthening the likelihood that this multimodal network analysis technique is yielding true positive results.

Age is a well-known factor in brain atrophy measures (Opfer et al., 2018), which we also observed. In fact, age has been found to affect the majority of grey matter atrophy measures typically associated with disease, e.g. in multiple sclerosis (Barkhof et al., 2009). Our study indicates that there may also be disease-related atrophy in the pallidum, cortical thickness and NBV in larger cohorts, when adjusted for age. Despite earlier studies finding no group differences with HC (Finke et al., 2016; Pasquier et al., 2019), these studies did not look at the complex interactions of MRI or OCT measures in relation to clinical attacks and disability measures, which multimodal network analysis allows for. Therefore, this study serves as an exemplary showcase for directed graph theory-based multimodal network analysis.

The cross-sectional exploratory nature and small sample size, including the low number of other attacks (brainstem or area postrema syndrome), are important limitations. To account for both physiological fluctuations and noisy data, we performed the multimodal network analysis using F-statistics calculated from an ANOVA with Type III sum of squares, including HC and patients. This method allows for accurate testing of main effects and interactions, as well as accounting for between-group and within-group variances (Kim, 2014). Inclusion of HC data and interaction effects in the models to minimize physiological and disease-unrelated effects is a clear strength of our approach. This method could be applied to studies where patients were treated with different attack preventing therapies to allow for evaluation of treatment–imaging relationships. We were not able to investigate this in our study due to the cross-sectional and retrospective nature, as well as subgrouping the heterogeneous attack preventing therapies would not have led to sufficient group sizes for any meaningful analysis. It is thus, a limitation that we cannot fully differentiate if there are central nervous system volume effects based on attack preventing therapy, neurodegeneration, astrocytopathy or inflammation. However, one may find it feasible to use therapies as interaction effects in future larger and longitudinal studies. We were also limited by the low number of attacks in our patient cohort, where very few patients had combined attacks; however, we performed a leave-one-out-cross-validation analysis of the multivariable linear regression models used for the nucleus accumbens and caudate nucleus volumes. This validation analysis gave low mean squared error values, thus, indicating that despite a skewness towards lower attack counts, the complex regression models fit quite well and can aid in the prediction of volumetric changes based on real-world combined attack and disease duration data.

In summary, using graph theory-based multimodal network analysis, we identified two DGM imaging biomarker candidates in AQP4-IgG+ NMOSD. Rather than studying the localized region of disease-related attacks, multimodal network analysis gives insights into connectivity and interactions between multiple regions in the central nervous system that can lead to knowledge gain of damage as part of a complex, dynamic system. The network approach provides an avenue for future investigations of potential biomarkers, or associated changes to monitor for clinical trials to treat many neurologically complex diseases in large and heterogeneous datasets.

Acknowledgements

We would like to thank all participants who volunteered for these observational studies and the study nurses, MRI technicians (S. Pikol and C. Kraut), OCT technicians and residents who aided in collecting data analyzed in this study.

Funding

Funding for this work was provided from the Deutsche Forschungsgemeinschaft DFG (EXC 257) that was awarded to F.P. and A.U.B. We also acknowledge support with publication fees from the DFG and the Open Access Publication Funds of Charité - Universitätsmedizin Berlin.

Competing interests

F.C.O. was employed by Nocturne, unrelated to this project. J.K. received conference registration fees from Biogen and financial research support from Krankheitsbezogenes Kompetenzzentrum Multiple Sklerose (KKNMS), not related to this work. M.S. reports no competing interests. K.R. received research support from Novartis, Merck Serono and German Ministry of Education and Research as well as speaking fees and travel grants from Bayer Healthcare, Biogen Idec, Merck Serono, sanofi-aventis/Genzyme, Teva Pharmaceuticals, Roche, Novartis and Guthy Jackson Charitable Foundation. J.B-S. received speaking fees and travel grants from Bayer Healthcare, Sanofi-Aventis/Genzyme, Biogen and Teva Pharmaceuticals. F.P. is on the study steering committee of Novartis OCTIMs study; received speaker honoraria and travel funding from Bayer, Novartis, Biogen, Teva, Sanofi-Aventis/Genzyme, Merck Serono, Alexion, Chugai Pharmaceutical, MedImmune.
and Shire; is an associate editor of Neurology®, Neuroimmunology & Neuroinflammation and an academic editor of PLoS One; consulted for Sanofi Genzyme, Biogen, MedImmune, Shire and Alexion; and received research support from Bayer, Novartis, Biogen, Teva, Sanofi-Aventis/Genzyme, Alexion, Merck Serono, the German Research Council, Werth Stiftung of the City of Cologne, the German Ministry of Education and Research, Arthur Arteins Stiftung Berlin, EU FP7 Framework Program, the Guthy Jackson Charitable Foundation, the Arthur Arteins Foundation and the National Multiple Sclerosis Society. A.U.B. is cofounder and shareholder of medical technology start-ups Nocturne GmbH and Motognosis GmbH. He is named as inventor on several patent applications describing multiple sclerosis serum biomarkers, retinal image analysis and visual perceptive computing for measuring motor dysfunction. The remaining authors declare no conflicts of interest.

References


